

Example:

$$X_i \sim N(\mu, \sigma^2)$$

$$P(U < c) = 0.9? \quad U = \frac{\sqrt{n}(\bar{X}_n - \mu)}{\sigma'}$$

$$\text{where } \sigma' = \sqrt{\frac{\sum_{i=1}^n (X_i - \bar{X}_n)^2}{n-1}}$$

$$\bar{X}_n \sim N(\mu, \sigma^2/n) \Rightarrow \frac{\sqrt{n}(\bar{X}_n - \mu)}{\sigma} \sim N(0, 1)$$

by theorem of joint dist. of \bar{X}_n and sample variance.

$$\sum_{i=1}^n \left(\frac{X_i - \bar{X}_n}{\sigma}\right)^2 \sim \chi^2_{(n-1)}$$

By definition of the t-distribution we know that if $Z \sim N(0, 1)$ and $Y \sim \chi^2_{(n)}$ and they are independent,

$$\text{then } U = \frac{Z}{\sqrt{\frac{Y}{n}}} \sim t_{(n)}$$

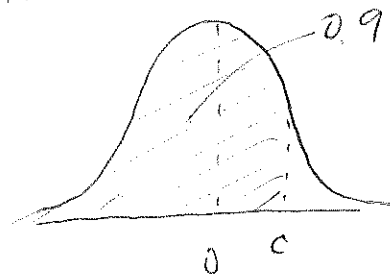
$$\text{So, considering } Z = \frac{\sqrt{n}(\bar{X}_n - \mu)}{\sigma} \text{ and } Y = \sum_{i=1}^n \left(\frac{X_i - \bar{X}_n}{\sigma}\right)^2$$

$$\text{we get that } \frac{\sqrt{n}(\bar{X}_n - \mu)}{\sqrt{\frac{\sum_{i=1}^n (X_i - \bar{X}_n)^2}{n-1}}} \sim t_{(n-1)}.$$

So, back to the example:

$$U = \frac{\sqrt{n}(\bar{X}_n - \mu)}{\sigma'} \sim t_{(n-1)}$$

$$\text{Now: } P(U < c) = 0.9$$



c is the 0.9 quantile of a t distributed random variable with $n-1$ degrees of freedom.

from a table of the t distribution with $n-1=10-1=9$ degrees of freedom or using $qt(0.9, df=9)$ we get that $c = 1.383$.

What can we say about $\bar{X}_n - \mu$?

$$P(U < 1.383) = 0.9$$

$$P\left(\frac{\sqrt{n}(\bar{X}_n - \mu)}{\sigma'} < 1.383\right) = 0.9$$

$$P\left(\bar{X}_n - \mu < \frac{1.383\sigma'}{\sqrt{n}}\right) = 0.9$$

$$P(\bar{X}_n - \mu < 0.437\sigma') = 0.9$$

Confidence intervals.

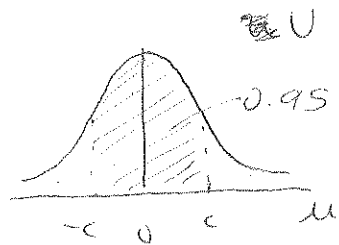
a) $X_i \stackrel{iid}{\sim} N(\mu, \sigma^2)$

b) \bar{X}_n is an estimator of μ

$\sigma' = \sqrt{\frac{\sum_{i=1}^n (X_i - \bar{X}_n)^2}{n-1}}$ is an estimator of σ^2 .

$\# U = \frac{\sqrt{n} (\bar{X}_n - \mu)}{\sigma'} \sim t_{(n-1)}$

c) $P(-c < U < c) = 0.95$



$$P(-c < U < c) = P(U < c) - P(U < -c)$$

$$= P(U < c) - P(U > c)$$

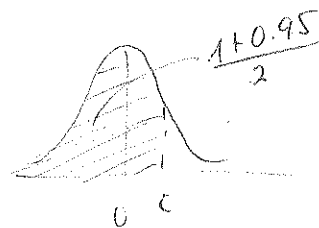
$$= P(U < c) - [1 - P(U < c)]$$

$$= \underbrace{2P(U < c)}_{T_{n-1}(c)} - 1$$

$$2T_{n-1}(c) - 1 = 0.95$$

$$T_{n-1}(c) = \frac{1+0.95}{2}$$

$$c = T_{n-1}^{-1}\left(\frac{1+0.95}{2}\right)$$



c is the $\frac{1+0.95}{2}$ quantile of a t distribution with $n-1$ degrees of freedom.

More generally, if we want to find c such that

$$P(-c < U < c) = \gamma, \text{ then } c = T_{n-1}^{-1}\left(\frac{1+\gamma}{2}\right), 0 < \gamma < 1$$

d) find ~~A and B~~ A and B such that $P(A < \mu < B) = 0.95$

From c)

$$P(-c < U < c) = 0.95, \quad c = T_{n-1}^{-1} \left(\frac{1+0.95}{2} \right)$$

$$P\left(-c < \frac{\sqrt{n}(\bar{X}_n - \mu)}{\sigma'} < c\right) = 0.95$$

$\longleftarrow U = \frac{\sqrt{n}(\bar{X}_n - \mu)}{\sigma'}$

$$P\left(-\frac{\sigma' c}{\sqrt{n}} < \bar{X}_n - \mu < \frac{\sigma' c}{\sqrt{n}}\right) = 0.95$$

$$P\left(-\bar{X}_n - \frac{c\sigma'}{\sqrt{n}} < -\mu < -\bar{X}_n + \frac{c\sigma'}{\sqrt{n}}\right) = 0.95$$

$$P\left(\bar{X}_n - \frac{c\sigma'}{\sqrt{n}} < \mu < \bar{X}_n + \frac{c\sigma'}{\sqrt{n}}\right) = 0.95$$

so $A = \bar{X}_n - \frac{c\sigma'}{\sqrt{n}}$

$B = \bar{X}_n + \frac{c\sigma'}{\sqrt{n}}$

$$c = T_{n-1}^{-1} \left(\frac{1+0.95}{2} \right)$$

$$\sigma' = \sqrt{\sum_{i=1}^n \left(\frac{x_i - \bar{X}_n}{\sigma} \right)^2}$$

More general $P(A < \mu < B) = \gamma, \quad 0 < \gamma < 1$

where $A = \bar{X}_n - \frac{c\sigma'}{\sqrt{n}}$

$B = \bar{X}_n + \frac{c\sigma'}{\sqrt{n}}$

$$c = T_{n-1}^{-1} \left(\frac{1+\gamma}{2} \right)$$

$$\sigma' = \sqrt{\sum_{i=1}^n \frac{(x_i - \bar{X}_n)^2}{n-1}}$$