

Suppose that we observe $X_1 = x_1$, and $X_2 = x_2$ sequentially, so

$$\xi(\theta | x_1, x_2) \propto f(x_2 | \theta) \xi(\theta | x_1)$$

the normalizing constant is given by

$$\frac{1}{\int_{\Omega} f(x_2 | \theta) \xi(\theta | x_1) d\theta}$$

$$\begin{aligned} \int_{\Omega} f(x_2 | \theta) \xi(\theta | x_1) d\theta &= \int_{\Omega} \frac{f(x_2 | \theta) f(x_1 | \theta) \xi(\theta)}{g(x_1)} d\theta \\ &= \int_{\Omega} \frac{f(x_2, x_1 | \theta) \xi(\theta)}{g(x_1)} d\theta \\ &= \int_{\Omega} \frac{h(x_2, x_1, \theta)}{g(x_1)} d\theta \\ &= \frac{g(x_1, x_2)}{g(x_1)} \\ &= h_1(x_2 | x_1) \end{aligned}$$

so, if we observe $x_1 = x_1, \dots, x_n = x_n$, then the normalizing constant would be

$$\frac{1}{\int_{\Omega} f(x_n | \theta) \xi(\theta | x_1, \dots, x_{n-1}) d\theta} = \frac{1}{g(x_n | x_1, \dots, x_{n-1})}$$

if we now observe $x_{n+1} = x_{n+1}$, we have that

$$\frac{1}{\int_{\Omega} f(x_{n+1} | \theta) \xi(\theta | x_1, \dots, x_n) d\theta} = \frac{1}{g(x_{n+1} | x_1, \dots, x_n)} \text{ so}$$

$$g(x_{n+1} | x_1, \dots, x_n) = \int_{\Omega} f(x_{n+1} | \theta) \xi(\theta | x_1, \dots, x_n) d\theta.$$

Example: lifetime of a new component.

We have $X_i | \theta \stackrel{iid}{\sim} \exp(\theta)$, $i=1, \dots, n$

$\theta | a, b \sim \text{Gamma}(a, b)$

We know that

$$X \sim \text{Gamma}(a, b) \rightarrow \frac{b^a}{\Gamma(a)} x^{a-1} e^{-bx}$$

$\theta | x_1, \dots, x_n \sim \text{Gamma}(a+n, b + \sum_{i=1}^n x_i)$, $n=5$

$$P(X_6 > 1.5 | x_1, \dots, x_5) = \int_{1.5}^{\infty} \underbrace{f(x_6 | x_1, \dots, x_5)}_{?} dx_6$$

$$f(x_6 | x_1, \dots, x_5) = \int_{\mathcal{R}} f(x_6 | \theta) \xi(\theta | x_1, \dots, x_5) d\theta$$

$$= \int_0^{\infty} \theta e^{-\theta x_6} \cdot \frac{(b + \sum x_i)^{a+n}}{\Gamma(a+n)} \theta^{a+n-1} e^{-(b + \sum x_i)\theta} d\theta$$

constant as a function of θ

$$= \frac{(b + \sum x_i)^{a+n}}{\Gamma(a+n)} \int_0^{\infty} \theta^{a+n} e^{-(b + \sum x_i + x_6)\theta} d\theta$$

Gamma($a+n+1$, $b + \sum_{i=1}^5 x_i + x_6$)

$$= \frac{(b + \sum x_i)^{a+n}}{\Gamma(a+n)} \cdot \frac{\Gamma(a+n+1)}{(b + \sum_{i=1}^5 x_i + x_6)^{a+n+1}}$$

$$\Gamma(n+1) = n\Gamma(n)$$

$$= (a+n) \frac{(b + \sum x_i)^{a+n}}{(b + \sum x_i + x_6)^{a+n+1}}$$

you can check that

$$\int \frac{(a+n)(b + \sum x_i)^{a+n}}{(b + \sum x_i + x_6)^{a+n+1}} dx_6 = 1$$

check it!

$$\int y^{-n} dy$$

$n=5$

Finally

$$P(X_6 > 1.5 | x_1, \dots, x_5) = \int_{1.5}^{\infty} \frac{(a+n)(b + \sum_{i=1}^5 x_i)^{a+n}}{(b + \sum x_i + x_6)^{a+n+1}} dx_6 = \dots = 0.1279$$

$$b) \mu_1 = \frac{\sigma^2 \mu_0 + n v_0^2 \bar{x}_n}{\sigma^2 + n v_0^2}$$

$$= \underbrace{\frac{\sigma^2}{\sigma^2 + n v_0^2}} \mu_0 + \underbrace{\frac{n v_0^2}{\sigma^2 + n v_0^2}} \bar{x}_n$$

• σ^2, v_0^2 are fixed, $n \rightarrow \infty$

• v_0^2, n fixed, $\sigma^2 \rightarrow \infty$

• σ^2, n fixed, $v_0^2 \rightarrow \infty$

$$c) P(\theta > 1 | \underline{x}), \quad \theta | \underline{x} \sim N(\mu_1, v_1^2)$$

$$P(\theta > 1 | \underline{x}) = 1 - P(\theta \leq 1 | \underline{x})$$

$$= 1 - P\left(z \leq \frac{1 - \mu_1}{\sqrt{v_1^2}}\right), \quad z \sim N(0, 1)$$

$$= 1 - \Phi\left(\frac{1 - \mu_1}{\sqrt{v_1^2}}\right)$$

↳ cumulative distr. function of z
(C.D.F.)