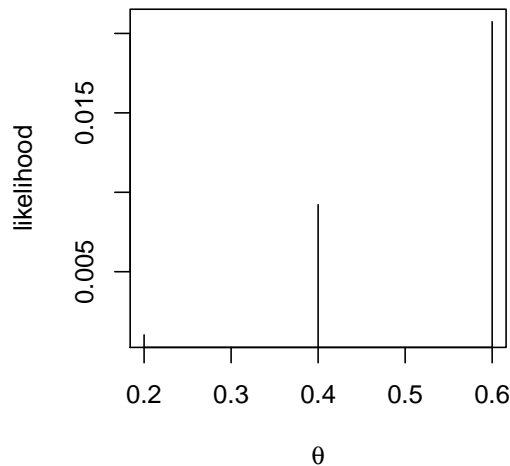


Exam 1: Class, Version B

Name:

1. Suppose that X_1, \dots, X_n are random variables describing whether a student approves its AMS 132 class or not. Assume that X_1, \dots, X_n form a random sample from the Bernoulli distribution with unknown probability θ , where θ can be 0.2, 0.4, or 0.6. A small random sample of 6 students was observed and 4 of them approved the class. **[15 pts]**

- (a) Plot the likelihood function. Specify the axes labels and values. **[5 pts]**



- (b) Find the maximum likelihood estimate of θ . **[5 pts]**

The maximum likelihood estimate is $\hat{\theta} = 0.6$

- (c) Find the maximum likelihood estimate of the variance of X_1 . **[5 pts]**

Notice that $Var(X_1) = \theta(1 - \theta) = g(\theta)$. Then, by the invariance property of the MLE we have that the maximum likelihood estimate of $Var(X_1)$ is $g(\hat{\theta}) = 0.6(1 - 0.6) = 0.24$

Name:

2. Suppose that X_1, \dots, X_4 are random variables describing the monthly amount of accumulated rainfall during winter in Santa Cruz. Assume that X_1, \dots, X_4 form a random sample from a gamma distribution with parameters $a > 0$ and $\theta > 0$, where a is known and θ is unknown. the density function is given by

$$f(x | a, \theta) = \frac{\theta^a}{\Gamma(a)} x^{a-1} e^{-\theta x}, \quad x > 0.$$

Consider $a = 2$ and that the total amount of rainfall during winter is 6, this is, $\sum_{i=1}^4 x_i = 6$. **[15 pts]**

- (a) Write the statistical model. **[5 pts]**

Let X be the random variable describing the monthly amount of accumulated rainfall during winter in Santa Cruz. The joint distribution of the random variables is

$$f(\mathbf{x} | \theta) = \prod_{i=1}^n \frac{\theta^a}{\Gamma(a)} x_i^{a-1} e^{-\theta x_i} = \frac{\theta^{na}}{\Gamma(a)^n} \prod_{i=1}^n x_i^{a-1} e^{-\theta \sum_{i=1}^n x_i},$$

where $\theta \in \Omega = \mathbb{R}^+$.

- (b) Consider an exponential prior distribution with mean equal to $1/3$ for θ . Considering square error loss function, find the Bayes estimate for θ . **[5 pts]**

Assuming that $\theta \sim \exp(\alpha)$, and noting that $E(\theta) = 1/3$, it follows that $\theta \sim \exp(3)$. First, lets find the posterior distribution of θ

$$\xi(\theta | \mathbf{x}) \propto \prod_{i=1}^n \frac{\theta^a}{\Gamma(a)} x_i^{a-1} e^{-\theta x_i} \xi(\theta) \propto \theta^{na} e^{-\theta \sum_{i=1}^n x_i} e^{-3\theta} \propto \theta^{na} e^{-\theta(3 + \sum_{i=1}^n x_i)}.$$

Therefore, $\theta | \mathbf{x} \sim \text{Gamma}(na+1, 3 + \sum_{i=1}^n x_i)$. Considering square error loss, Bayes estimate is the posterior mean. So, Bayes estimate is given by

$$\delta^* = \frac{na + 1}{3 + \sum_{i=1}^n x_i} = \frac{9}{9} = 1.$$

- (c) Under the same prior distribution for θ as in (b), and considering square error loss function, find the Bayes estimate for the variance of the monthly amount of accumulated rainfall during winter in Santa Cruz. **[5 pts]**

The variance of the monthly amount of accumulated rainfall during winter in Santa Cruz is given by $\text{Var}(X) = \frac{a}{\theta^2} = \frac{2}{\theta^2}$. Here we want to estimate parameter $\Psi = \frac{2}{\theta^2}$. Considering square error loss, Bayes estimate is the posterior mean of Ψ . So,

$$\begin{aligned} \delta^* &= E\left(\frac{2}{\theta^2} | \mathbf{x}\right) = \int_0^\infty \frac{2}{\theta^2} \frac{9^9}{\Gamma(9)} \theta^8 e^{-9\theta} d\theta = \frac{2 * 9^9}{\Gamma(9)} \int_0^\infty \theta^6 e^{-9\theta} d\theta, \\ &= \frac{2 * 9^9 \Gamma(7)}{\Gamma(9) 9^7} = \frac{3 * 9^2}{7 * 8}. \end{aligned}$$

Exam 1: Take home, Version B

Due date is Wednesday, February 7, at 4:00 pm, at Merrill Acad 102 (before the class starts).

Late take home part will not be accepted!

1. Suppose that X_1, \dots, X_n form a random sample from a Poisson distribution with unknown parameter θ . [15 pts]

- (a) Write the coefficient of variation of X , denoted $CV(X)$, as a function of the parameter θ . [3 pts]

Hint: Recall that $CV(X) = \frac{\sqrt{Var(X)}}{|E(X)|}$. Since $E(X) = Var(X) = \theta$, X_i for a random sample, and $CV(X) = \frac{\sqrt{Var(X)}}{|E(X)|}$, it follows that $CV(X) = \frac{\sqrt{\theta}}{|\theta|} = 1/\sqrt{\theta}$.

- (b) Find the maximum likelihood estimate of θ . [4 pts]

The likelihood function of θ is

$$f_n(\mathbf{x} | \theta) = \frac{e^{-n\theta} \theta^{\sum_{i=1}^n x_i}}{\prod_{i=1}^n x_i!}.$$

Its logarithm is

$$L(\theta) = -n\theta + \sum_{i=1}^n x_i \log(\theta) - \log(1/\prod_{i=1}^n x_i!).$$

Its derivative is

$$\frac{d}{d\theta} L(\theta) = -n + \frac{1}{\theta} \sum_{i=1}^n x_i.$$

Making the derivative equal to 0 and solving the equation we get that a candidate for being the maximum likelihood estimate of θ is

$$\hat{\theta} = \bar{X}_n = \frac{1}{n} \sum_{i=1}^n x_i.$$

Finally, checking that

$$\frac{d^2}{d\theta^2} L(\theta) = -\frac{1}{\theta^2} \sum_{i=1}^n x_i < 0,$$

for every sample and θ , we conclude that $\hat{\theta}$ is the maximum likelihood estimate of θ .

- (c) Find the maximum likelihood estimate of the coefficient of variation of X , denoted $CV(X)$. [4 pts]

From part (a) it follows that $CV(X) = 1/\sqrt{\theta} = g(\theta)$. By the invariance property of the M.L.E., it follows that the maximum likelihood estimate of $CV(X)$ is $g(\hat{\theta}) = 1/\sqrt{\hat{\theta}} = 1/\sqrt{\bar{X}_n}$.

- (d) Assume now that θ is a random parameter and has a gamma prior distribution with parameters $a > 0$ and $b > 0$. Consider square error loss function and find the Bayes estimate for θ . Under what kind of prior distribution are the maximum likelihood estimate and the Bayes estimate

equals? **[4 pts]**

Since θ has a Gamma prior distribution and X_i follows a Poisson distribution, by theorem from class, it follows that

$$\theta \mid \mathbf{x} \sim \text{Gamma}\left(a + \sum_{i=1}^n x_i, b + n\right).$$

Under square error loss function, Bayes estimate, δ^* , is the posterior mean, therefore

$$\delta^* = \frac{a + \sum_{i=1}^n x_i}{b + n}.$$

Finally, if $a = b = 0$, Bayes estimate and the maximum likelihood estimate are equal. This is, Bayes estimate and the maximum likelihood estimate are equal when an improper prior distribution for θ is considered.

2. Suppose that X_1, \dots, X_n form a random sample from a distribution that has p.d.f. given by

$$f(x \mid \theta) = \theta x e^{-0.5\theta x^2}, \quad x > 0, \quad \theta > 0.$$

Assume that θ is a random parameter with gamma prior distribution with parameters $a > 0$ and $b > 0$.

[15 pts]

(a) Show that the gamma prior distribution is a conjugate prior for samples from the distribution that has p.d.f. as given above. Explain with your own words what this means. **[7 pts]**

Note that

$$\xi(\theta \mid \mathbf{x}) \propto \theta^n e^{-0.5 \sum_{i=1}^n x_i^2} \theta^{a-1} e^{-b\theta} \propto \theta^{a+n-1} e^{-(0.5 \sum_{i=1}^n x_i^2 + b)\theta},$$

therefore, $\theta \mid \mathbf{x} \sim \text{Gamma}\left(a + n, 0.5 \sum_{i=1}^n x_i^2 + b\right)$.

This means, that if the parameter has a gamma prior distribution and the random variables have p.d.f. has defined above, then the posterior distribution of θ has also a gamma distribution.

(b) Find the p.d.f. of a new observation, X_{n+1} , given that $X_1 = x_1, \dots, X_n = x_n$ have been observed, this is, find $f(x_{n+1} \mid x_1, \dots, x_n)$. **[8 pts]**

$$\begin{aligned} f(x_{n+1} \mid \mathbf{x}) &= \int_0^\infty \theta x_{n+1} e^{-0.5\theta x_{n+1}^2} \frac{(0.5 \sum_{i=1}^n x_i^2 + b)^{a+n}}{\Gamma(a+n)} \theta^{a+n-1} e^{-(0.5 \sum_{i=1}^n x_i^2 + b)\theta} d\theta, \\ &= x_{n+1} \frac{(0.5 \sum_{i=1}^n x_i^2 + b)^{a+n}}{\Gamma(a+n)} \int_0^\infty \theta^{a+n+1-1} e^{-(0.5[\sum_{i=1}^n x_i^2 + x_{n+1}^2] + b)\theta} d\theta, \\ &= x_{n+1} \frac{(0.5 \sum_{i=1}^n x_i^2 + b)^{a+n}}{\Gamma(a+n)} \frac{\Gamma(a+n+1)}{(0.5[\sum_{i=1}^n x_i^2 + x_{n+1}^2] + b)^{a+n+1}}, \\ &= x_{n+1} (a+n) \frac{(0.5 \sum_{i=1}^n x_i^2 + b)^{a+n}}{(0.5[\sum_{i=1}^n x_i^2 + x_{n+1}^2] + b)^{a+n+1}}, \end{aligned}$$

$x_{n+1} > 0$.