

Examples: improper prior distributions:

a) we will consider $X_i | \theta \stackrel{i.i.d.}{\sim} \text{Bernoulli}(\theta)$, $(i=1, \dots, n)$
 $\theta \sim \text{Beta}(\alpha, \beta)$ $\Rightarrow \xi(\theta) = \frac{\Gamma(\alpha+\beta)}{\Gamma(\alpha)\Gamma(\beta)} \theta^{\alpha-1} (1-\theta)^{\beta-1}$ $p(x_i|\theta) = \theta^{x_i} (1-\theta)^{1-x_i}$
 prior hyperparameters

so, the posterior distribution is:

$$\xi(\theta | x_1, \dots, x_n) \propto f_n(x | \theta) \xi(\theta)$$

$$\propto \theta^{\sum_{i=1}^n x_i} (1-\theta)^{n - \sum_{i=1}^n x_i} \theta^{\alpha-1} (1-\theta)^{\beta-1}$$

$$\propto \theta^{\alpha + \sum x_i - 1} (1-\theta)^{\beta + n - \sum x_i - 1}$$

so $\theta | \underline{x} \sim \text{Beta}(\alpha + \sum_{i=1}^n x_i, \beta + n - \sum_{i=1}^n x_i)$

what happens if $\alpha = \beta = 0$.

$$\xi(\theta) \propto \theta^{-1} (1-\theta)^{-1} = \frac{1}{\theta} \cdot \frac{1}{1-\theta} = \frac{1}{\theta} + \frac{1}{1-\theta}$$

$$\int_0^1 \xi(\theta) d\theta = \int_0^1 \frac{1}{\theta} + \frac{1}{1-\theta} d\theta = \int_0^1 \frac{1}{\theta} d\theta + \int_0^1 \frac{1}{1-\theta} d\theta$$

$$= \log(\theta) \Big|_{\theta=0}^1 + [-\log(1-\theta)] \Big|_{\theta=0}^1 = \infty.$$

check it!
(log = ln)

in this case we have that

~~$\theta | \underline{x} \sim \text{Beta}(\sum x_i, n - \sum x_i)$~~ $\theta | \underline{x} \sim \text{Beta}(\sum_{i=1}^n x_i, n - \sum_{i=1}^n x_i)$

if $n=1$ and $\sum_{i=1}^n x_i = x_i = 0$, then the posterior is improper. $\int_0^1 \xi(\theta | x) d\theta = \infty$.

if $n > 1$, $\sum_{i=1}^n x_i = 0$, then the posterior distribution is $\xi(\theta | \underline{x}) \propto \theta^{-1} (1-\theta)^n$

$$\int_0^1 \frac{(1-\theta)^n}{\theta} d\theta = \text{HW. check if it is finite or not.}$$

is not finite!

$$b) X_i | \theta \sim \text{Poisson}(\theta) \rightarrow f(x_i | \theta) = \frac{e^{-\theta} \theta^{x_i}}{x_i!}$$

$$\theta \sim \text{Gamma}(\alpha, \beta) \rightarrow \xi(\theta) = \frac{\beta^\alpha}{\Gamma(\alpha)} \theta^{\alpha-1} e^{-\beta\theta}$$

$$\xi(\theta | \underline{x}) \propto e^{-n\theta} \theta^{\sum x_i} \theta^{\alpha-1} e^{-\beta\theta}$$

$$= \theta^{\alpha + \sum x_i - 1} e^{-(\beta+n)\theta}$$

$$\text{so } \theta | \underline{x} \sim \text{Gamma}\left(\alpha + \sum_{i=1}^n x_i, \beta + n\right)$$

If $\alpha = \beta = 0$.

$$\xi(\theta) \propto \frac{1}{\theta}$$

$$\int_{\Omega} \xi(\theta) d\theta = \int_0^{\infty} \frac{1}{\theta} d\theta = \ln \theta \Big|_0^{\infty} = \infty$$

then $\theta | \underline{x} \sim \text{Gamma}(\sum x_i, n)$

If nobody arrives to the store in an hour, then $\sum x_i = n = 0$ and the posterior distr. is improper.

c) $X_i | \theta, \sigma^2 \sim N(\theta, \sigma^2)$, σ^2 is known

$$\theta \sim N(\mu_0, \nu_0^2) \rightarrow \xi(\theta) = \frac{1}{\sqrt{2\pi\nu_0^2}} \exp\left\{-\frac{1}{2\nu_0^2}(\theta - \mu_0)^2\right\}$$

what happens if $\nu_0^2 \rightarrow \infty$?

$$\xi(\theta) \propto \exp\left\{-\frac{1}{2\nu_0^2}(\theta - \mu_0)^2\right\}$$

as $\nu_0^2 \rightarrow \infty$ the prior distr. of θ , $\xi(\theta)$, is a constant.

$$\text{and } \int_{-\infty}^{\infty} \xi(\theta) d\theta = \infty$$